

# Supplementary Materials of SC<sup>2</sup>-PCR

Zhi Chen<sup>1</sup>

Kun Sun<sup>2</sup>

Fan Yang<sup>1</sup>

Wenbing Tao<sup>1\*</sup>

<sup>1</sup>National Key Laboratory of Science and Technology on Multi-spectral Information Processing, School of Artificial Intelligence and Automation, Huazhong University of Science and Technology, China

<sup>2</sup>Hubei Key Laboratory of Intelligent Geo-Information Processing, School of Computer Science, China University of Geosciences, China  
{z.chen, fanyang, wenbingtao}@hust.edu.cn; sunkun@cug.edu.cn

## 1. Derivations

### 1.1. Derivation of ambiguity probability

**Notations.** Suppose there are  $N$  pairs of correspondences, in which the inlier rate is  $\alpha$ . The second order spatial compatibility  $SC^2$  measure between correspondence  $i$  and  $j$  is defined as follows:

$$SC_{ij}^2 = C_{ij} \cdot \sum_{k=1}^N C_{ik} \cdot C_{kj}. \quad (1)$$

$C$  ( $C \in \mathbb{R}^{N \times N}$ ) is the hard compatibility matrix as follows:

$$C_{ij} = \begin{cases} 1; d_{ij} \leq d_{thr} \\ 0; d_{ij} > d_{thr} \end{cases}. \quad (2)$$

$d_{ij}$  is the distance difference between two correspondences. When  $i = j$ ,  $C_{ij}$  is set to 0. The distance difference between two inliers  $d_{in,in}$  follows the uniform distribution, and the probability density function of  $d_{in,in}$  is as follows:

$$\text{PDF}_{in,in}(l) = 1/d_{thr}, 0 \leq l \leq d_{thr} \quad (3)$$

For the distance difference between an inlier and an outlier  $d_{in,out}$ , or two outliers  $d_{out,out}$ , their probability density functions are both  $F(\cdot)$ :

$$\text{PDF}_{in,out}(l) = F(l), \text{PDF}_{out,out}(l) = F(l); 0 \leq l \leq d_r. \quad (4)$$

$F(l)$  is a constant within  $(0, d_{thr})$  as follows:

$$F(l) = f_0, 0 \leq l \leq d_{thr}. \quad (5)$$

**Remark 1.1:** The ambiguity probability of  $SC^2$  measure, i.e.,  $P(SC_{in,out}^2 > SC_{in,in}^2)$ , can be written as follows:

$$\begin{aligned} P(SC_{in,out}^2 > SC_{in,in}^2) &= p \cdot P(X > (N \cdot \alpha - 2)), \\ X &\sim S((N\alpha - 1)p + (N(1 - \alpha) - 1)p^2, N(1 - \alpha)p^2), \\ p &= d_{thr} \cdot f_0. \end{aligned} \quad (6)$$

**Derivation of Remark 1.1.** We first reformulate Eq. (1) as follows:

$$SC_{ij}^2 = C_{ij} \cdot M_{ij}, \quad (7)$$

where  $M_{ij}$  is computed as follows:

$$M_{ij} = \sum_{k=1}^N C_{ik} \cdot C_{kj}. \quad (8)$$

$M_{ij}$  counts the quantity of the commonly compatible correspondences of  $i$  and  $j$  in the global set. According to Eq. (2) and (3), we can obtain that:

$$P(C_{in,in} = 1) = 1. \quad (9)$$

According to Eq. (2), (4) and (5), we can get that

$$P(C_{in,out} = 1) = \int_0^{d_{thr}} F(l)dl = d_{thr} \cdot f_0 = p \quad (10)$$

$$P(C_{out,out} = 1) = \int_0^{d_{thr}} F(l)dl = d_{thr} \cdot f_0 = p. \quad (11)$$

According to Eq. (7), to make  $SC_{in,out}^2 > SC_{in,in}^2$  hold, two conditions need to be met:  $C_{in,out} = 1$  and  $M_{in,out} > M_{in,in}$ . According to Eq. (10), we can obtain the following equation:

$$\begin{aligned} &P(SC_{in,out}^2 > SC_{in,in}^2) \\ &= P(C_{in,out} = 1) \cdot P(M_{in,out} > M_{in,in}) \\ &= p \cdot P(M_{in,out} > M_{in,in}). \end{aligned} \quad (12)$$

Next, we compute the distribution of  $M_{in,out}$  and  $M_{in,in}$ . Since inliers have different distribution with outliers, we consider them separately and reformulate Eq. 8 as follows:

$$M_{ij} = \sum_{m \in \mathcal{I}} C_{im} \cdot C_{mj} + \sum_{n \in \mathcal{O}} C_{in} \cdot C_{nj}. \quad (13)$$

\*Corresponding author.

where  $\mathcal{I}$  is the inlier set while  $\mathcal{O}$  is the outlier set. (For convenience we use this notation in the following part).

We first discuss the value in  $M$  matrix between two inliers, i.e.  $M_{in,in}$ . According to Eq. (9), we can find that any two inliers are compatible. Thus, when correspondence  $i$  and  $j$  are inliers, the number of correspondences compatible with both of them in the inlier set is the number of inliers excluding themselves ( $C_{ii} = 0, C_{jj} = 0$ ), i.e.:

$$\sum_{m \in \mathcal{I}} C_{im} \cdot C_{mj} = N \cdot \alpha - 2; i \in \mathcal{I}, j \in \mathcal{I}, \quad (14)$$

where  $\alpha$  is the inlier rate. For outliers, according to Eq. (10), the probability that an outlier is compatible with an inlier is  $p$ . Then the probability that an outlier is compatible with both  $i$  and  $j$  is  $p^2$ . The number of outliers in the whole correspondence set is  $N(1 - \alpha)$ . So the number of correspondences compatible with both of them in the outlier set is in a Bernoulli distribution [2] as follows:

$$\sum_{n \in \mathcal{O}} C_{in} \cdot C_{nj} \sim B(N(1 - \alpha), p^2); i \in \mathcal{I}, j \in \mathcal{I}, \quad (15)$$

where  $B(\cdot, \cdot)$  is the Bernoulli distribution. Thus,  $M_{in,in}$  is in the following distribution:

$$M_{in,in} \sim N \cdot \alpha - 2 + B(N(1 - \alpha), p^2). \quad (16)$$

After that, we discuss the distribution of the value in  $M$  matrix between an inlier and an outlier, i.e.  $M_{in,out}$ . For convenience, we assume correspondence  $i$  is inlier while  $j$  is outlier. For the inlier set except correspondence  $i$  ( $C_{ii} = 0$ ), any of them is compatible with  $i$  (Eq. 9), and the probability that one of them is compatible with  $j$  is  $p$  (Eq. 10). So the number of correspondences compatible with both of correspondence  $i$  and  $j$  in the inlier set is in following distribution:

$$\sum_{m \in \mathcal{I}} C_{im} \cdot C_{mj} \sim B(N\alpha - 1, p); i \in \mathcal{I}, j \in \mathcal{O}. \quad (17)$$

Meanwhile, for each outlier except correspondence  $j$  ( $C_{jj} = 0$ ), the probabilities that it is compatible with  $i$  or  $j$  are both  $p$  (Eq. 10 and 11). So the probability that an outlier is both compatible with  $i$  and  $j$  is  $p^2$ . Thus, we can get the following distribution:

$$\sum_{n \in \mathcal{O}} C_{in} \cdot C_{nj} \sim B(N(1 - \alpha) - 1, p^2); i \in \mathcal{I}, j \in \mathcal{O}. \quad (18)$$

So the distribution of  $M_{in,out}$  is as follows:

$$M_{in,out} \sim B(N\alpha - 1, p) + B(N(1 - \alpha) - 1, p^2); \quad (19)$$

Since  $p$  is a small value, the Binomial distribution in Eq. (16) and (19) can be approximately equivalent to the Poisson distribution [2], i.e.:

$$\begin{aligned} M_{in,in} &\sim N \cdot \alpha - 2 + \pi(N(1 - \alpha)p^2), \\ M_{in,out} &\sim \pi((N\alpha - 1)p) + \pi((N(1 - \alpha) - 1)p^2), \end{aligned} \quad (20)$$

where  $\pi(\cdot)$  is the Poisson distribution. Furthermore, for two Poisson distribution:  $X_1 \sim \pi(\lambda_1)$  and  $X_2 \sim \pi(\lambda_2)$ , their sum is also in the Poisson distribution [2] as follows:

$$X_1 + X_2 \sim \pi(\lambda_1 + \lambda_2). \quad (21)$$

So we can convert  $M_{in,out}$  in Eq. (20) into following form:

$$M_{in,out} \sim \pi((N\alpha - 1)p + (N(1 - \alpha) - 1)p^2). \quad (22)$$

Meanwhile, we can convert  $P(M_{in,out} > M_{in,in})$  into following form:

$$\begin{aligned} &P(M_{in,out} > M_{in,in}) \\ &= P(M_{in,out} - M_{in,in} > 0) \\ &= P(X > N \cdot \alpha - 2), \end{aligned} \quad (23)$$

where  $X$  is in following distribution:

$$X \sim \pi((N\alpha - 1)p + (N(1 - \alpha) - 1)p^2) - \pi(N(1 - \alpha)p^2) \quad (24)$$

For two Poisson distribution:  $X_1 \sim \pi(\lambda_1)$  and  $X_2 \sim \pi(\lambda_2)$ , their difference is in the Skellam distribution [5–7], i.e.:

$$X_1 - X_2 \sim S(\lambda_1, \lambda_2). \quad (25)$$

So the distribution of  $X$  in Eq. (24) can be converted as follows:

$$S((N\alpha - 1)p + (N(1 - \alpha) - 1)p^2, N(1 - \alpha)p^2). \quad (26)$$

Combining Eq. (12), (23) and (26), we compute the value of  $P(SC_{in,out}^2 > SC_{in,in}^2)$  as Remark 1.1.

## 2. Additional Experiments

**Parameter  $K_1$  and  $K_2$ .** In the proposed method, when some seeds are selected, we use a two-stage selection strategy to extend each seed into a consensus set. In the first stage,  $K_1$  correspondences are selected by finding top- $K_1$  neighbors of seed. In the second stage, a local  $SC^2$  matrix is rebuilt to further filter potential outliers and reserve  $K_2$  correspondences. In order to show the effect of these two parameters, we report the registration results with respect to these parameters in Tab. 1.

As shown in Tab. 1, our method is parameter insensitive. The proposed method with different parameters can all lead to acceptable results combined with both FPFH and FCGF descriptors. In the final version, we choose  $K_1=30$  and  $K_2=20$  for its best registration recall.

**Heat Map.** Compared with the spatial compatibility (SC) measure [1, 8], the second order spatial compatibility ( $SC^2$ ) measure can reduce the probability of ambiguity event. In order to show the difference between these two measures, we report the heat maps of them on a real data in

Param		FPFH			FCGF			Time(s)
$K_1$	$K_2$	RR(%)	RE(o)	TE(cm)	RR(%)	RE(o)	TE(cm)	
10	3	82.13	2.15	6.67	92.91	2.05	6.52	0.11
10	5	82.69	2.10	6.59	93.04	2.05	6.52	0.11
20	10	83.79	2.13	6.56	93.10	2.06	6.54	0.11
30	20	83.98	2.18	6.56	93.28	2.08	6.55	0.11
40	30	83.67	2.15	6.71	93.22	2.05	6.52	0.12
50	40	83.67	2.16	6.78	93.16	2.06	6.54	0.12
60	50	83.92	2.18	6.68	93.10	2.05	6.49	0.12
70	60	83.55	2.16	6.69	92.98	2.04	6.49	0.13

Table 1. The registration results with varying parameters on 3DMatch dataset.

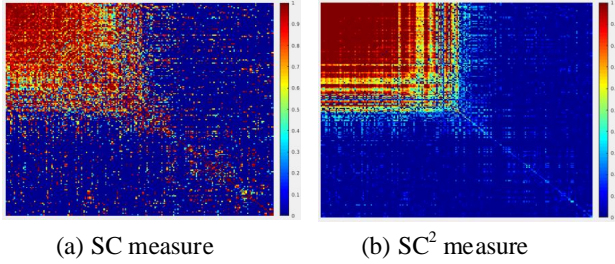


Figure 1. The heat maps of spatial compatibility (SC) and second order spatial compatibility (SC<sup>2</sup>) measures.

Fig. 1. Specifically, we sort all the correspondences, placing inliers first and outliers behind to build the SC and SC<sup>2</sup> matrices respectively. Then we normalize these two matrices to [0,1], and plot the corresponding heat maps. As shown in Fig. 1, there are high values on the right and bottom in the heat map of SC measure, showing that many outliers are compatible with inliers. By contrast, the right and bottom sides are clean in the heat map of SC<sup>2</sup> measure, showing that outliers have low compatibility with inliers.

**Scene-wise Results.** Following [1, 3], we also report the scene-wise registration results of the proposed method on 3DMatch dataset, combining both FPFH and FCGF descriptors as shown in Tab. 2.

	FPFH			FCGF		
	RR(%)	RE(o)	TE(cm)	RR(%)	RE(o)	TE(cm)
Kitchen	88.34	1.95	5.41	99.21	1.69	5.15
Home1	89.74	1.82	6.23	96.79	1.79	6.37
Home2	73.56	2.80	7.46	83.17	3.48	7.50
Hotel1	92.04	2.20	7.01	98.67	1.89	6.08
Hotel2	81.73	2.08	6.49	91.35	1.94	5.61
Hotel3	90.74	2.01	5.79	92.59	2.08	5.80
Study	76.03	2.31	8.94	88.36	2.31	9.21
Lab	76.62	1.67	5.99	80.52	1.91	8.44

Table 2. Scene-wise registration results on 3DMatch Dataset.

**More Qualitative Results.** We show the more registration results in Fig. 2 and 3.

## References

- [1] Xuyang Bai, Zixin Luo, Lei Zhou, Hongkai Chen, Lei Li, Zeyu Hu, Hongbo Fu, and Chiew-Lan Tai. Pointdsc: Robust point cloud registration using deep spatial consistency. In *Proceedings of the IEEE/CVF Conference on Computer Vision and Pattern Recognition*, pages 15859–15869, 2021. 2, 3, 4
- [2] George EP Box, William H Hunter, Stuart Hunter, et al. *Statistics for experimenters*, volume 664. John Wiley and sons New York, 1978. 2
- [3] Christopher Choy, Wei Dong, and Vladlen Koltun. Deep global registration. In *Proceedings of the IEEE/CVF Conference on Computer Vision and Pattern Recognition*, pages 2514–2523, 2020. 3
- [4] Martin A Fischler and Robert C Bolles. Random sample consensus: a paradigm for model fitting with applications to image analysis and automated cartography. *Communications of the ACM*, 24(6):381–395, 1981. 4
- [5] Joseph Oscar Irwin. The frequency distribution of the difference between two independent variates following the same poisson distribution. *Journal of the Royal Statistical Society*, 100(3):415–416, 1937. 2
- [6] Dimitris Karlis and Ioannis Ntzoufras. Analysis of sports data by using bivariate poisson models. *Journal of the Royal Statistical Society: Series D (The Statistician)*, 52(3):381–393, 2003. 2
- [7] Dimitris Karlis and Ioannis Ntzoufras. Bayesian analysis of the differences of count data. *Statistics in medicine*, 25(11):1885–1905, 2006. 2
- [8] Siwen Quan and Jiaqi Yang. Compatibility-guided sampling consensus for 3-d point cloud registration. *IEEE Transactions on Geoscience and Remote Sensing*, 58(10):7380–7392, 2020. 2
- [9] Qian-Yi Zhou, Jaesik Park, and Vladlen Koltun. Fast global registration. In *European Conference on Computer Vision*, pages 766–782. Springer, 2016. 4



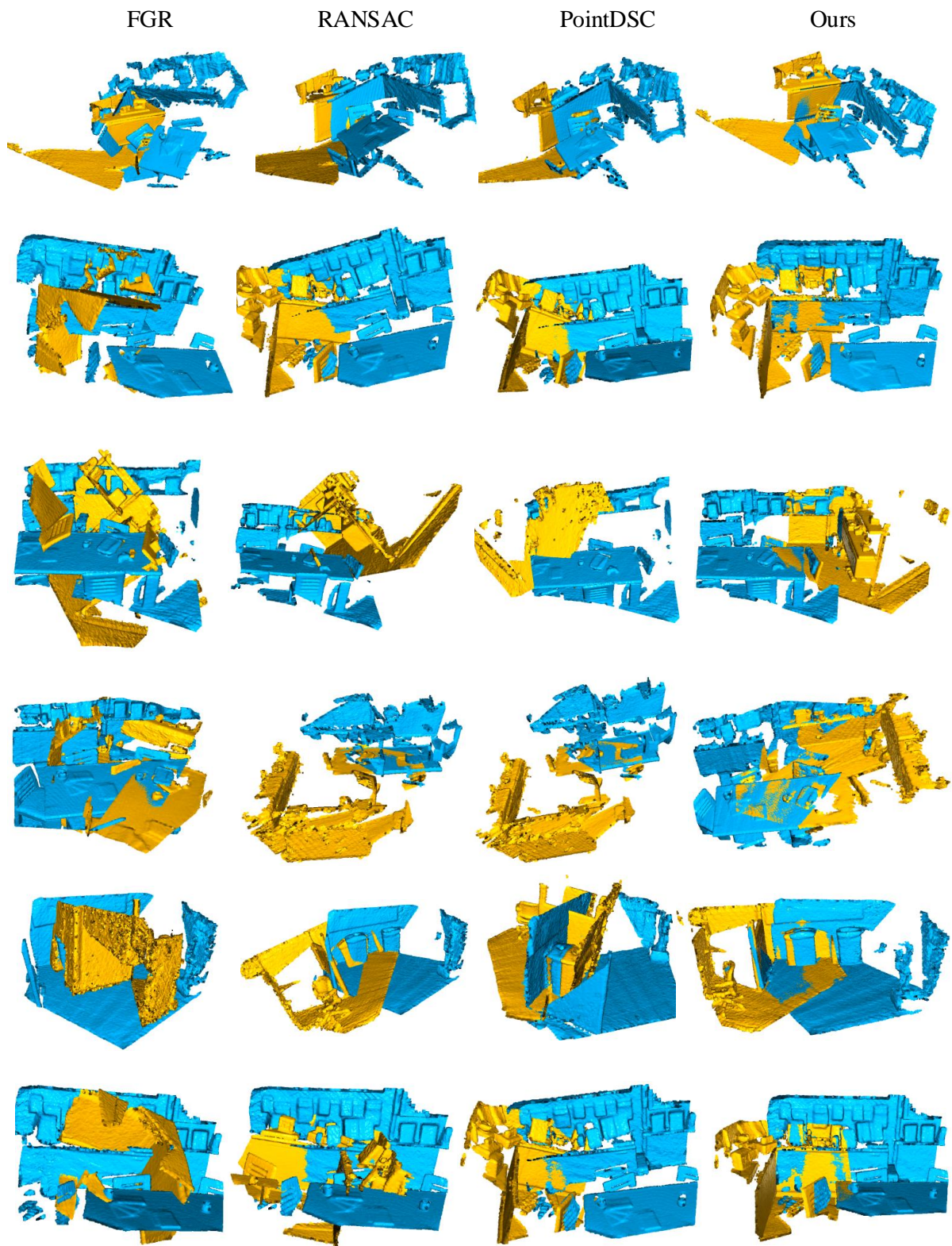


Figure 2. Qualitative comparison on 3DMatch and 3DLoMatch dataset. From left to right are: FGR [9], RANSAC [4], PointDSC [1] and Ours

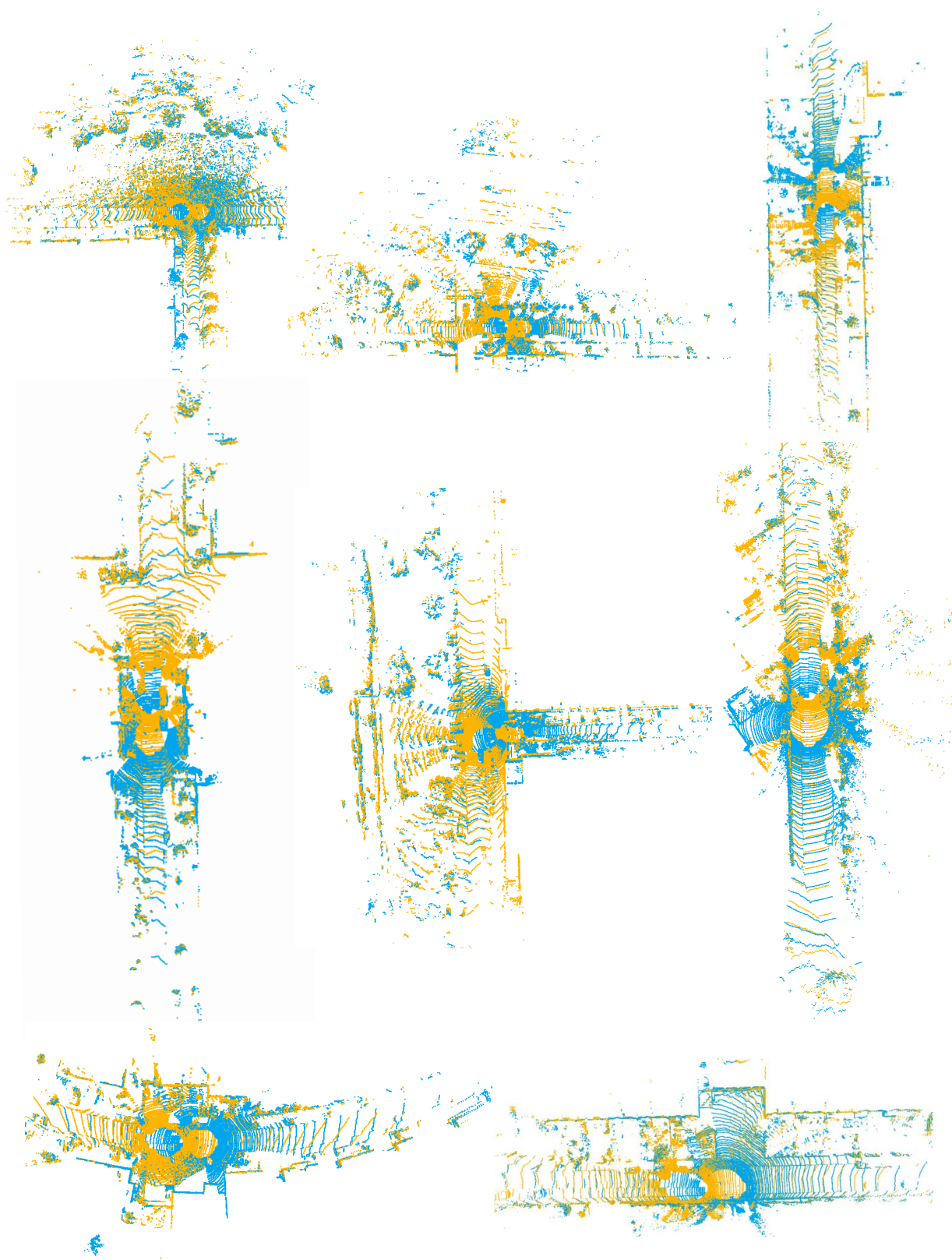


Figure 3. Qualitative results of our method on KITTI dataset.