

Subspace Constraint and Contribution Estimation for Heterogeneous Federated Learning

Supplementary Material

6. Detailed Configuration

In addition to the hyperparameter settings outlined in the main paper, we follow the configurations recommended in the original papers for each baseline method. SOLO requires no additional hyperparameters. For FML [39], we set the knowledge distillation parameters as $\alpha = 0.5$ and $\beta = 0.5$. For FedKD [48], we match the auxiliary model learning rate to the client learning rate, both set at 0.01, with $T_{\text{start}} = 0.95$ and $T_{\text{end}} = 0.95$. For FedDistill [19], we use $\gamma = 0.1$. For FedProto [42], we set $\lambda = 0.1$. For FedTGP [57], we configure $\lambda = 0.1$ (the same as FedProto), $\tau = 100$, and $S = 100$ (server training epochs). The feature vector dimension is set to 512, as suggested in [57].

7. Label Distribution Skew

Fig. 8 illustrates the data distribution across varying levels of heterogeneity in the *Label skew* scenario, with the number of clients set to 20. We use the size of the point to represent the sample size.

8. Effect of Subspace Update Interval s

In Tab. 7, we present the impact of the subspace update interval on model performance. Our observations indicate that FedSCE outperforms the optimal baseline across all update intervals, demonstrating its robustness to hyperparameters. Additionally, accuracy can be further enhanced by carefully tuning the subspace update interval.

	<i>Office10</i>				<i>Cifar10</i>			
s	1	3	5	10	1	3	5	10
Acc	75.55	75.70	76.34	76.66	72.90	72.99	72.86	72.85

Table 7. **Hyperparameter Study:** Test accuracy (%) of FedSCE under different s for the *Office10* and *Cifar10* datasets. Please refer to Sec. 8 for a detailed discussion.

9. Further Ablation Experiments

As shown in Tab. 8, we conducted ablation experiments on the adjustment strategies on the client (LSL) and server (CE) to verify their effectiveness. In addition, we further show the mean and standard deviation of different components under three different random seeds in Tab. 9.

10. Discussion and Limitation

The performance of a model fundamentally depends on its generalization ability and convergence behavior. Fed-

LSL	CE	<i>Amaz</i>	<i>Calt</i>	<i>DSLR</i>	<i>Webc</i>	<i>C.Avg</i>	<i>D.Avg</i>
✓		82.04	63.34	77.21	90.54	78.29	74.45
	✓	83.29	63.70	78.48	91.89	79.34	75.31
		84.34	64.59	79.74	92.56	80.31	76.26
✓	✓	84.55	64.59	79.74	92.57	80.36	76.34

Table 8. **Ablation** experiment with adjustments on the client (LSL) and server (CE). Please see details in Sec. 9.

LSL	DML	CE	<i>C.Avg</i>	<i>D.Avg</i>
✓			78.49±0.13	73.29±0.10
✓			79.53±0.28	74.08±0.16
✓	✓		79.65±0.25	75.45±0.31
✓	✓	✓	80.40±0.10	76.47±0.14

Table 9. The **mean** and **standard deviation** of the performance of different components. Please see details in Sec. 9.

SCE enhances generalization performance by constraining model updates to a low-rank subspace and reducing the complexity of the model space. While traditional dimensionality reduction techniques may improve generalization, they usually affect the convergence speed. Exploiting the inherent low-rank property of the update trajectory, FedSCE adopts dynamic singular value decomposition (SVD) to continuously refine the subspace, effectively minimizing the projection error while maintaining convergence. However, if we only use a fixed subspace that is randomly created at the start, it often leads to much larger projection errors, which can slow down or prevent proper convergence.

In addition, considering the computational and storage overheads associated with subspace updates, we provide the per-round computation time and average client memory utilization of various algorithms on the *Office10* dataset in Tab. 10. FedSCE incurs a manageable level of memory and computation overhead due to subspace updates, whereas FedKD [48] requires more storage and computation resources due to local SVD calculations.

	SOLO	FML	KD	Distill	Proto	TGP	Our
M.cost	0.18	0.22	0.31	0.18	0.18	0.18	0.24
Time	20.01	27.82	45.53	30.50	29.62	32.17	30.43

Table 10. **Memory cost** (GB) and **running time** (second) of different algorithms. Please see details in Sec. 10.

11. Generalization Analysis for Subspace Constraint

Let $\mathcal{X} \subset \mathbb{R}^D$ represent the input space, $\mathcal{Z} \subset \mathbb{R}^I$ the latent feature space, and $\mathcal{Y} \subset \mathbb{R}$ the output space. A feature mapping function $\mathcal{G} : \mathcal{X} \mapsto \mathcal{Z}$ projects inputs into a feature space. We define a domain \mathcal{T} as a combination of a data distribution \mathcal{D} over \mathcal{X} and a labeling function $c^* : \mathcal{X} \mapsto \mathcal{Y}$. Given a domain $\mathcal{T} := \langle \mathcal{D}, c^* \rangle$ and a representation \mathcal{G} , let $\tilde{\mathcal{D}}$ denote the induced distribution of \mathcal{D} under \mathcal{G} [4], such that for any event \mathcal{B} ,

$$\mathbb{E}_{z \sim \tilde{\mathcal{D}}}[\mathcal{B}(z)] = \mathbb{E}_{x \sim \mathcal{D}}[\mathcal{B}(\mathcal{G}(x))]. \quad (17)$$

Let $h : \mathcal{Z} \mapsto \mathcal{Y}$ denote a hypothesis that maps features to labels, and let $\mathcal{H} \subseteq \{h : \mathcal{Z} \mapsto \mathcal{Y}\}$ be a hypothesis class. For two distributions \mathcal{D} and \mathcal{D}' , the \mathcal{H} -divergence $d_{\mathcal{H}}(\mathcal{D}, \mathcal{D}')$ between \mathcal{D} and \mathcal{D}' is defined as:

$$d_{\mathcal{H}}(\mathcal{D}, \mathcal{D}') := 2 \sup_{\mathcal{A} \in \mathcal{A}_{\mathcal{H}}} |\Pr_{\mathcal{D}}(\mathcal{A}) - \Pr_{\mathcal{D}'}(\mathcal{A})|, \quad (18)$$

where $\mathcal{A}_{\mathcal{H}}$ is the set of measurable subsets induced by $h \in \mathcal{H}$. Additionally, $\mathcal{H}\Delta\mathcal{H}$ represents the *symmetric difference hypothesis space* [5]:

$$\mathcal{H}\Delta\mathcal{H} := \{h(z) \oplus h'(z) \mid h, h' \in \mathcal{H}\}, \quad (19)$$

where \oplus denotes the XOR operation, indicating disagreement between $h(z)$ and $h'(z)$. Let $\mathcal{A}_{\mathcal{H}\Delta\mathcal{H}}$ be the corresponding set of measurable subsets. The *distribution divergence* $d_{\mathcal{H}\Delta\mathcal{H}}(\cdot, \cdot)$ induced by the symmetric difference hypothesis space is given by [5]:

$$d_{\mathcal{H}\Delta\mathcal{H}}(\mathcal{D}, \mathcal{D}') := 2 \sup_{\mathcal{A} \in \mathcal{A}_{\mathcal{H}\Delta\mathcal{H}}} |\Pr_{\mathcal{D}}(\mathcal{A}) - \Pr_{\mathcal{D}'}(\mathcal{A})|. \quad (20)$$

Now, let \mathcal{D} and \mathcal{D}' be two distributions over the input space \mathcal{X} , and let $\tilde{\mathcal{D}}$ and $\tilde{\mathcal{D}}'$ denote their respective induced distributions over \mathcal{G} . Then, using the definition of $d_{\mathcal{H}\Delta\mathcal{H}}(\cdot, \cdot)$, we have:

$$\begin{aligned} d_{\mathcal{H}\Delta\mathcal{H}}(\tilde{\mathcal{D}}, \tilde{\mathcal{D}}') &= 2 \sup_{\mathcal{A} \in \mathcal{A}_{\mathcal{H}\Delta\mathcal{H}}} |\mathbb{E}_{x \sim \mathcal{D}}[\Pr(\mathcal{A}(\mathcal{G}(x)))] - \mathbb{E}_{x \sim \mathcal{D}'}[\Pr(\mathcal{A}(\mathcal{G}(x)))]| \\ &= 2 \sup_{\mathcal{A} \in \mathcal{A}_{\mathcal{H}\Delta\mathcal{H}}} |\mathbb{E}_{z \sim \tilde{\mathcal{D}}}[\Pr(\mathcal{A}(z))] - \mathbb{E}_{z \sim \tilde{\mathcal{D}}'}[\Pr(\mathcal{A}(z))]| \\ &= 2 \sup_{\mathcal{A} \in \mathcal{A}_{\mathcal{H}\Delta\mathcal{H}}} |\Pr_{\tilde{\mathcal{D}}}(\mathcal{A}) - \Pr_{\tilde{\mathcal{D}}'}(\mathcal{A})|. \end{aligned} \quad (21)$$

Then, we provide the theorem for the generalization performance from [4, 5, 56, 61] under their assumptions.

Theorem 1 Let \mathcal{T}_S and \mathcal{T}_T be the source and target domains, whose data distributions are \mathcal{D}_S and \mathcal{D}_T . Let $\mathcal{G} : \mathcal{X} \mapsto \mathcal{Z}$ be a feature representation function, and $\tilde{\mathcal{D}}_S, \tilde{\mathcal{D}}_T$ be the induced images of \mathcal{D}_S and \mathcal{D}_T over \mathcal{G} , respectively. Let \mathcal{H} be a set of hypotheses with VC-dimension d . Then with probability at least $1 - \delta$, $\forall h \in \mathcal{H}$:

$$\mathcal{L}_{\mathcal{T}_T}(h) \leq \hat{\mathcal{L}}_{\mathcal{T}_S}(h) + \sqrt{\frac{4}{m} \left(d \log \frac{2em}{d} + \log \frac{4}{\delta} \right)} + d_{\mathcal{H}\Delta\mathcal{H}}(\tilde{\mathcal{D}}_S, \tilde{\mathcal{D}}_T) + \lambda, \quad (22)$$

where e is the base of the natural logarithm, $\hat{\mathcal{L}}_{\mathcal{T}_S}(h)$ is the empirical risk of the source domain given m observable samples, and $\lambda = \min_{h \in \mathcal{H}}(\mathcal{L}_{\mathcal{T}_T}(h) + \mathcal{L}_{\mathcal{T}_S}(h))$ is the optimal risk on the two domains.

In the context of hierarchical federated learning (HFL), where the emphasis is on transferring knowledge from global to local models on the client side [56], we can restate Theorem 1 as follows:

Theorem 2 Consider a virtual global data domain \mathcal{T} and a local data domain \mathcal{T}_m . Define $\mathcal{T} = \langle \mathcal{D}, c^* \rangle$ with $|\tilde{\mathcal{D}}| = n$ and $\mathcal{T}_m = \langle \mathcal{D}_m, c^* \rangle$, where $\mathcal{D}, \mathcal{D}_m \subseteq \mathcal{X}$, and $c^* : \mathcal{X} \mapsto \mathcal{Y}$ is a ground-truth labeling function. Let $\mathcal{G} : \mathcal{X} \mapsto \mathcal{Z}$ be a feature extraction function. Suppose \mathcal{H} represents a hypothesis space constrained by subspace with VC dimension d' , and $h : \mathcal{Z} \mapsto \mathcal{Y}$ for each $h \in \mathcal{H}$. Then, with probability at least $1 - \delta$, the following holds for any $h \in \mathcal{H}$:

$$\mathcal{L}_{\mathcal{T}_m}(h) \leq \hat{\mathcal{L}}_{\mathcal{T}}(h) + \sqrt{\frac{4}{n} \left(d' \log \frac{2en}{d'} + \log \frac{4}{\delta} \right)} + d_{\mathcal{H}\Delta\mathcal{H}}(\tilde{\mathcal{D}}, \tilde{\mathcal{D}}_m) + \lambda_m, \quad (23)$$

where $d' < d < 2n$, d denotes the VC dimension of the standard hypothesis space. $\lambda_m := \min_h(\mathcal{L}_{\mathcal{T}_m}(h) + \mathcal{L}_{\mathcal{T}}(h))$ denotes an oracle performance, e is the base of the natural logarithm, $\tilde{\mathcal{D}}, \tilde{\mathcal{D}}_m$ are the induced distributions of $\mathcal{D}, \mathcal{D}_m$ under \mathcal{G} . $d_{\mathcal{H}\Delta\mathcal{H}}(\tilde{\mathcal{D}}_m, \tilde{\mathcal{D}})$ denotes the divergence measured over a symmetric-difference hypothesis space [5]. $\hat{\mathcal{L}}_{\mathcal{T}}(h)$ is the empirical risk on \mathcal{T} .

12. Convergence Analysis in Non-convex Settings

In $\mathcal{L}_m^t, \mathcal{L}_m^{t,0}$, we use t represents the time step before the subspace update, and $t, 0$ represents the time step between the subspace update and the first iteration of the current round. In addition, we use t, e to denote the e -th local epoch of the t -th communication round. Next, we prove the non-convex convergence of the local cross entropy loss with dynamic subspace constraint.

Assumption 1 Each local objective function is L_1 -Lipschitz smooth; that is, $\|\nabla \mathcal{L}_m^{t_1,e} - \nabla \mathcal{L}_m^{t_2,e}\|_2 \leq L_1 \|\theta_m^{t_1,e} - \theta_m^{t_2,e}\|_2$, which implies the following quadratic bound, $\mathcal{L}_m^{t_1,e} - \mathcal{L}_m^{t_2,e} \leq \langle \nabla \mathcal{L}_m^{t_2,e}, (\theta_m^{t_1,e} - \theta_m^{t_2,e}) \rangle + \frac{L_1}{2} \|\theta_m^{t_1,e} - \theta_m^{t_2,e}\|_2^2, \forall t_1, t_2 > 0$.

Assumption 2 The stochastic gradient $g_m^{t,e} = \nabla \mathcal{L}(\theta_m^{t,e}, \xi_m^t)$ is an unbiased estimator of the local gradient for each client. $\mathbb{E}_{\xi_m \sim D_m}[g_m^{t,e}] = \nabla \mathcal{L}(\theta_m^{t,e}) = \nabla \mathcal{L}_m^{t,e}$, and its variance is bounded by σ^2 : $\mathbb{E}[\|g_m^{t,e} - \nabla \mathcal{L}(\theta_m^{t,e})\|_2^2] \leq \sigma^2$.

Assumption 3 The expectation of the stochastic gradient is bounded by G : $\mathbb{E}[\|g_m^{t,e}\|_2^2] \leq G^2, \forall t, e$.

Proposition 1 For the local update $\Delta \tilde{\theta}_m^{t,e}$, the squared projection error is bounded; that is, $\mathbb{E}[\|\text{Proj}(\Delta \tilde{\theta}_m^{t,e}) - \Delta \tilde{\theta}_m^{t,e}\|_2^2] \leq \rho_R \eta^2 e^2 G^2, \rho_R = \frac{\sum_{r=R+1}^K \sigma_r^2}{\sum_{r=1}^K \sigma_r^2}, \forall t, e$. σ_r denotes singular value.

In fact, we make this proposition based on the low-rank nature of the model update space. This may be a bit strict, but if we replace $\rho_R = \frac{\sum_{r=R+1}^K \sigma_r^2}{\sum_{r=1}^K \sigma_r^2}$ with $0 < \rho_R < 1$, the following convergence proof also holds.

12.1. Key Lemmas

Lemma 1 Let Assumption 1 and Assumption 3 hold. The total client loss of an arbitrary client can be bounded:

$$\mathbb{E}[\mathcal{L}_m^{t+1}] \leq \mathcal{L}_m^{t,0} - (\eta - \frac{L_1 \eta^2}{2}) \sum_{e=0}^{E-1} \|\nabla \mathcal{L}_m^{t,e}\|_2^2 + \frac{L_1 E \eta^2}{2} \sigma^2. \quad (24)$$

Proof. This lemma only focuses on local training at the client level, combined with the initial local goal. By replacing the relevant symbols in Lemma 1 of FedProto [42], it can be obtained.

Lemma 2 Let Assumption 1 to 3 and Proposition 1 hold. After the local subspace update, the loss function of an arbitrary client can be bounded as:

$$\mathbb{E}[\mathcal{L}_m^{t+1,0}] \leq \mathcal{L}_m^{t+1} + \mu \rho_R \eta^2 E^2 G^2. \quad (25)$$

Proof.

$$\begin{aligned} \mathcal{L}_m^{t+1,0} &= \mathcal{L}_m^{t+1} + \mathcal{L}_m^{t+1,0} - \mathcal{L}_m^{t+1} \\ &= \mathcal{L}_m^{t+1} + \mu \left\| P_t P_t^T \Delta \tilde{\theta}_m^{t,E} - \Delta \tilde{\theta}_m^{t,E} \right\|^2 - \mu \left\| P_{t+1} P_{t+1}^T \Delta \tilde{\theta}_m^{t,E} - \Delta \tilde{\theta}_m^{t,E} \right\|^2 \\ &\leq \mathcal{L}_m^{t+1} + \mu \left\| P_t P_t^T \Delta \tilde{\theta}_m^{t,E} - \Delta \tilde{\theta}_m^{t,E} \right\|^2 \\ &\leq \mathcal{L}_m^{t+1} + \mu \rho_R \eta^2 E^2 G^2. \end{aligned} \quad (26)$$

Take expectations of random variable ξ on both sides, then

$$\mathbb{E}[\mathcal{L}_m^{t+1,0}] \leq \mathcal{L}_m^{t+1} + \mu \rho_R \eta^2 E^2 G^2. \quad (27)$$

12.2. Theorems

Theorem 3 (One-round deviation) Let Assumption 1 to 3 and Proposition 1 hold. For an arbitrary client, after every communication round, we have,

$$\mathbb{E}[\mathcal{L}_m^{t+1,0}] \leq \mathcal{L}_m^{t,0} - (\eta - \frac{L_1 \eta^2}{2}) \sum_{e=0}^{E-1} \|\nabla \mathcal{L}_m^{t,e}\|_2^2 + \frac{L_1 E \eta^2}{2} \sigma^2 + \mu \rho_R \eta^2 E^2 G^2. \quad (28)$$

Proof. Taking the expectation of θ_m^t on both sides of Lemma 2, and then adding Lemma 1 and Lemma 2, we get Eq. (28).

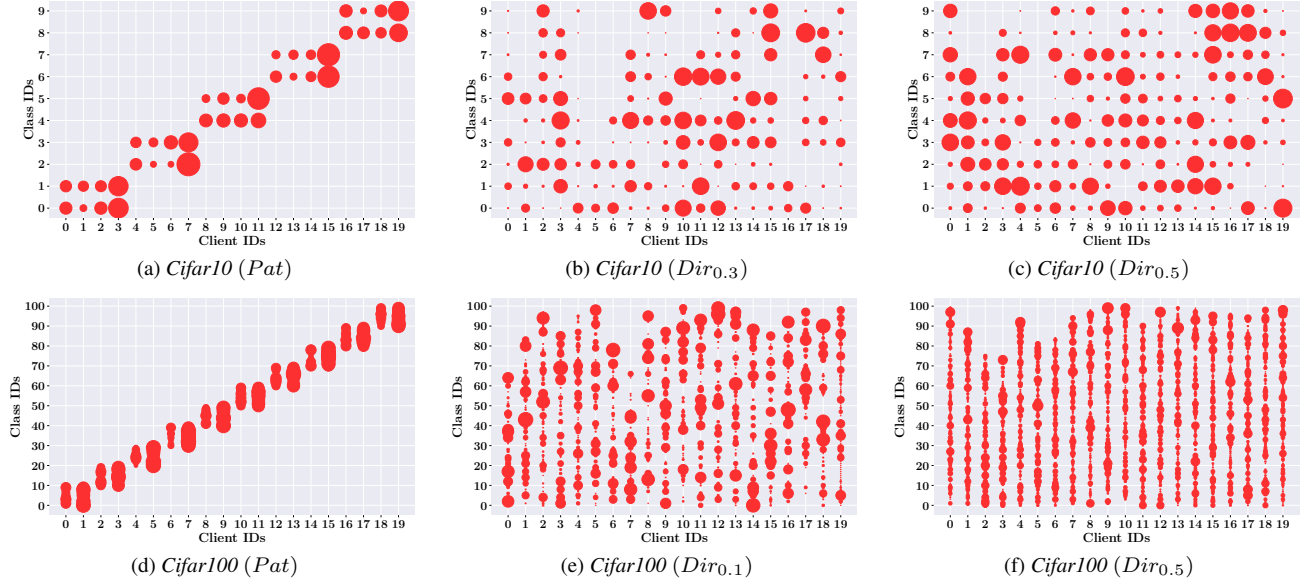


Figure 8. Data distribution of each client on *Cifar10* and *Cifar100* under different settings. The size of the circle indicates the number of samples.

Theorem 4 Assuming that Assumption 1 to 3 and Proposition 1 hold, \mathcal{L}_m^* denotes the local optimum, the following inequality is valid for any arbitrary client and any $\epsilon > 0$:

$$\frac{1}{TE} \sum_{t=0}^{T-1} \sum_{e=0}^{E-1} \mathbb{E}[\|\nabla \mathcal{L}_m^{t,e}\|_2^2] \leq \frac{L_1 \eta \sigma^2 + 2\mu \rho_R \eta E G^2}{2 - L_1 \eta} + \frac{2(\mathcal{L}_m^0 - \mathcal{L}_m^*)}{TE \eta (2 - L_1 \eta)} < \epsilon. \quad (29)$$

Proof. Considering the communication rounds of Eq. (28) from $t = 0$ to $t = T - 1$ and the local epoch of each communication round of Eq. (28) from $e = 0$ to $e = E$, we have

$$\frac{1}{TE} \sum_{t=0}^{T-1} \sum_{e=0}^{E-1} \mathbb{E}[\|\nabla \mathcal{L}_m^{t,e}\|_2^2] \leq \frac{\frac{1}{TE} \sum_{t=0}^{T-1} (\mathcal{L}_m^{t,0} - \mathbb{E}[\mathcal{L}_m^{t+1,0}]) + \frac{L_1 \eta^2}{2} \sigma^2 + \mu \rho_R \eta^2 E G^2}{\eta - \frac{L_1 \eta^2}{2}}. \quad (30)$$

Given any $\epsilon > 0$, let

$$\frac{\frac{1}{TE} \sum_{t=0}^{T-1} (\mathcal{L}_m^{t,0} - \mathbb{E}[\mathcal{L}_m^{t+1,0}]) + \frac{L_1 \eta^2}{2} \sigma^2 + \mu \rho_R \eta^2 E G^2}{\eta - \frac{L_1 \eta^2}{2}} < \epsilon, \quad (31)$$

Let $\Delta = \mathcal{L}_m^0 - \mathcal{L}_m^*$. Since $\sum_{t=0}^{T-1} (\mathcal{L}_m^{t,0} - \mathbb{E}[\mathcal{L}_m^{t+1,0}]) \leq \Delta$, the above equation holds when

$$\frac{\frac{2\Delta}{TE} + L_1 \eta^2 \sigma^2 + 2\mu \rho_R \eta^2 E G^2}{2\eta - L_1 \eta^2} = \frac{2(\mathcal{L}_m^0 - \mathcal{L}_m^*)}{TE \eta (2 - L_1 \eta)} + \frac{L_1 \eta \sigma^2 + 2\mu \rho_R \eta E G^2}{2 - L_1 \eta} < \epsilon, \quad (32)$$

that is:

$$T > \frac{2\Delta}{E\epsilon(2\eta - L_1 \eta^2) - E\eta(L_1 \eta \sigma^2 + 2\mu \rho_R \eta E G^2)}, \quad (33)$$

then, we have:

$$\frac{1}{TE} \sum_{t=0}^{T-1} \sum_{e=0}^{E-1} \mathbb{E}[\|\nabla \mathcal{L}_m^{t,e}\|_2^2] < \epsilon, \quad (34)$$

when

$$\eta < \frac{2\epsilon}{L_1(\epsilon + \sigma^2) + 2\mu \rho_R E G^2}. \quad (35)$$

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