

Supplementary of “Deep Generative Model for Robust Imbalance Classification”

Proof of Theorem 1 and 2

In this supplementary, we will theoretically analyze the generalization error bounds of the proposed model. Motivated by [3], we can define the generalization of input feature and label generation process by measuring the difference between the population real data distribution (\mathcal{P}_{real}) and the corresponding generated data distribution (\mathcal{P}_G). The generalization error will be acceptable if this population distance is close to the empirical distance between the observed real data distribution ($\tilde{\mathcal{P}}_{real}$) and the corresponding generated data distribution ($\tilde{\mathcal{P}}_G$). In the proposed deep latent variable model, given latent code Z , the input feature information X is independent to input label Y , i.e., ($X \perp Y|Z$). In this case, the data distribution (X, Y) $\sim \mathcal{P}_{real}$ can be factorized into two parts $X \sim \mathcal{X}_{real}$ and $Y \sim \mathcal{Y}_{real}$, similar to \mathcal{P}_G , $\tilde{\mathcal{P}}_{real}$ and $\tilde{\mathcal{P}}_G$.

Definition 2. For the empirical real distribution ($\tilde{\mathcal{X}}_{real}, \tilde{\mathcal{Y}}_{real}$) with N training examples, a generated distribution ($\tilde{\mathcal{X}}_G, \tilde{\mathcal{Y}}_G$) generalizes under the distribution distance $d(\cdot, \cdot)$ with generalization error $\delta_1, \delta_2 > 0$ if the following holds with high probability,

$$|E(X) - E(\tilde{X})| \leq \delta_1 \quad (8)$$

$$|E(Y) - E(\tilde{Y})| \leq \delta_2 \quad (9)$$

where

$$E(X) = \mathbb{E}_{\mathbf{x}^{(real)} \sim \mathcal{X}_{real}, \mathbf{x}^{(G)} \sim \mathcal{X}_G} d(\mathbf{x}^{(real)}, \mathbf{x}^{(G)}),$$

$$E(\tilde{X}) = \frac{1}{N} \sum_{i=1}^N d(\tilde{\mathbf{x}}_i^{(real)}, \tilde{\mathbf{x}}_i^{(G)}) | \tilde{\mathbf{x}}_i^{(real)} \sim \tilde{\mathcal{X}}_{real}, \tilde{\mathbf{x}}_i^{(G)} \sim \tilde{\mathcal{X}}_G$$

$$E(Y) = \mathbb{E}_{\mathbf{y}^{(real)} \sim \mathcal{Y}_{real}, \mathbf{y}^{(G)} \sim \mathcal{Y}_G} d(\mathbf{y}^{(real)}, \mathbf{y}^{(G)}),$$

$$E(\tilde{Y}) = \frac{1}{N} \sum_{i=1}^N d(\tilde{\mathbf{y}}_i^{(real)}, \tilde{\mathbf{y}}_i^{(G)}) | \tilde{\mathbf{y}}_i^{(real)} \sim \tilde{\mathcal{Y}}_{real}, \tilde{\mathbf{y}}_i^{(G)} \sim \tilde{\mathcal{Y}}_G$$

$E(\mathcal{X})$ and $E(\mathcal{Y})$ indicate the population distance between the real and generated distributions on feature and label information respectively. $E(\tilde{\mathcal{X}})$ and $E(\tilde{\mathcal{Y}})$ are the corresponding empirical distances.

Theorem 3. For any $X \in \mathbb{R}^{D \times N}$ ($N, D > 0$),

$$E(X) \leq E(\tilde{X}) + \sqrt{\frac{\log \delta_1}{-2N} (\max_i d_i)^2}$$

holds with probability at least $1 - \delta_1$ ($\delta_1 > 0$) over uniformly choosing an empirical version (\tilde{X}) of X . Here, $d_i = d(\tilde{\mathbf{x}}_i^{(real)}, \tilde{\mathbf{x}}_i^{(G)}) = \|\tilde{\mathbf{x}}_i^{(real)} - \tilde{\mathbf{x}}_i^{(G)}\|_2^2$.

Proof 1. Since the entries in X are chosen independently and uniformly, it is reasonable to assume each $d_i = d(\tilde{\mathbf{x}}_i^{(real)}, \tilde{\mathbf{x}}_i^{(G)})$ is a random variable and satisfies $p(\zeta \geq d_i \geq 0) = 1$, where $\zeta = \max_i d_i$. Hence, based on the Hoeffding Inequality, we have $p(|E(X) - E(\tilde{X})| \geq \epsilon) \leq \exp(-\frac{2N\epsilon^2}{\zeta^2})$. By setting $\epsilon = \sqrt{\frac{\log \delta_1}{-2N} \zeta^2}$, we have $p(|E(X) - E(\tilde{X})| \leq \sqrt{\frac{\log \delta_1}{-2N} \zeta^2}) \geq 1 - \delta_1$, i.e.,

$$p\left(E(\mathbf{X}) \leq E(\tilde{\mathbf{X}}) + \sqrt{\frac{\log \delta_1}{-2N} (\max_i d_i)^2}\right) \geq 1 - \delta_1$$

Therefore, the error of generative model for feature information is bounded.

Theorem 4. Given prior label probabilities $\{p_1, \dots, p_c, \dots, p_C\}$ (where $p_c = P(y = c)$ and $\sum_{c=1}^C p_c = 1$) and the conditional latent variable densities $\{f_1, f_2, \dots, f_C\}$ (where $f_c = f(\mathbf{z}|y = c)$), following [30], the error rate of generative classifier can be formulated and bounded:

$$\epsilon^C = 1 - \int \max\{p_1 f_1(\mathbf{z}), \dots, p_C f_C(\mathbf{z})\} d\mathbf{z} \leq \delta_2$$

Proof 2. According to the definition of ϵ^C , we can get

$$\epsilon^C = \mathbb{E}_{\mathbf{z}} [1 - \max_{c=1, \dots, C} p(y = c|\mathbf{z})] \quad (10)$$

Denote $a_u = p(y = u|\mathbf{z})$ and they are sorted in an ascending order, i.e. $a_C = \max\{a_u|_{u=1}\}$, we have $\sum_{v=1}^C a_v = 1$ and

$$\begin{aligned} 1 - a_C &= \sum_{u=1}^{C-1} a_u = \sum_{u=1}^{C-1} \sum_{v=1}^C a_u a_v \\ &= 2 \sum_{u=1}^{C-2} \sum_{v=u+1}^{C-1} a_u a_v + \sum_{u=1}^{C-1} a_u a_C + \sum_{u=1}^{C-1} a_u^2 \\ &\leq 2 \sum_{u=1}^{C-2} \sum_{v=u+1}^{C-1} a_u a_v + \sum_{u=1}^{C-1} a_u a_C + \sum_{u=1}^{C-1} a_u a_C \end{aligned} \quad (11)$$

Let denote the marginal distribution of Z as

$$f^{(C)}(\mathbf{z}) = \sum_{c=1}^C p_c f_c(\mathbf{z}) = \sum_{c=1}^C p_c f(\mathbf{z}|y = c)$$

$$\delta_{u,v}^C := \int \frac{p_u f_u(\mathbf{z}) p_v f_v(\mathbf{z})}{f^{(C)}(\mathbf{z})} d\mathbf{z} = \mathbb{E}_{\mathbf{z}} [a_u a_v]$$

According to the definition of Expectation, we know

$$\begin{aligned} \delta_{u,v}^C &= \mathbb{E}_{\mathbf{z}} [a_u a_v] = 2 \sum_{u=1}^{C-1} \sum_{v=u+1}^C a_u a_v \\ &= 2 \sum_{u=1}^{C-2} \sum_{v=u+1}^{C-1} a_u a_v + 2 \sum_{u=1}^{C-1} a_u a_C. \end{aligned} \quad (12)$$

Obviously, $1 - a_C \leq \delta_{u,v}^C$. By setting $\delta_2 = \delta_{u,v}^C$, $\epsilon^C \leq \delta_2$ holds.